

Written Exam at the Department of Economics winter 2019-20

**Financial Econometrics A**

Final Exam

January 14, 2020

(3-hour closed book exam)

Answers only in English.

**This exam question consists of 5 pages in total**

**Falling ill during the exam**

If you fall ill during an examination at Peter Bangs Vej, you must:

- contact an invigilator who will show you how to register and submit a blank exam paper.
- leave the examination.
- contact your GP and submit a medical report to the Faculty of Social Sciences no later than five (5) days from the date of the exam.

**Be careful not to cheat at exams!**

You cheat at an exam, if during the exam, you:

- Make use of exam aids that are not allowed
- Communicate with or otherwise receive help from other people
- Copy other people's texts without making use of quotation marks and source referencing, so that it may appear to be your own text
- Use the ideas or thoughts of others without making use of source referencing, so it may appear to be your own idea or your thoughts
- Or if you otherwise violate the rules that apply to the exam

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Please note there is a total of **8** questions that you should provide answers to. That is, **3** questions under *Question A*, and **5** under *Question B*.

## Question A:

Consider the model for  $x_t \in \mathbb{R}$  given by

$$x_t = \sigma_t z_t, \quad \sigma_t^2 = \omega + \alpha x_{t-1}^2$$

with  $(z_t)$  an *i.i.d.* process, and  $z_t$  is scaled  $t$ -distributed and given by

$$z_t = \sqrt{\frac{v-2}{v}} \tau_t,$$

where  $\tau_t$  is  $t$ -distributed with  $v = 5$  degrees of freedom. In particular, we have

$$E\tau_t = 0, \quad V(\tau_t) = E(\tau_t^2) = \frac{v}{v-2} = \frac{5}{3} \quad \text{and} \quad E(\tau_t^4) = \frac{3v^2}{(v-4)(v-2)} = 25.$$

Moreover, the model parameters satisfy that  $\omega > 0$  and  $\alpha \geq 0$ .

**Question A.1:** Show that  $x_t$  is weakly mixing with  $E x_t^4 < \infty$  if  $\alpha < 1/3$ .

**Question A.2:** Consider the extended ARCH(2) model,

$$x_t = \sigma_t z_t, \quad \sigma_t^2 = \omega + \alpha x_{t-1}^2 + \beta x_{t-2}^2,$$

with  $\omega > 0$  and  $\alpha, \beta \geq 0$ , and  $z_t$  still *i.i.d.* scaled  $t$ -distributed with  $v = 5$ . The Gaussian QMLE  $\hat{\theta} = (\hat{\omega}, \hat{\alpha}, \hat{\beta})'$  is obtained by maximizing  $L(\theta)$  given by

$$L(\theta) = -\frac{1}{T} \sum_{t=1}^T \left( \log \sigma_t^2(\theta) + \frac{x_t^2}{\sigma_t^2(\theta)} \right),$$
$$\sigma_t^2(\theta) = \omega + \alpha x_{t-1}^2 + \beta x_{t-2}^2.$$

It follows that,

$$S_\beta(\theta) = \partial L(\theta) / \partial \beta = \frac{1}{T} \sum_{t=1}^T \left( \frac{x_t^2}{\sigma_t^2(\theta)} - 1 \right) \frac{x_{t-2}^2}{\sigma_t^2(\theta)}.$$

With  $\theta_0 = (\omega_0, \alpha_0, 0)$  (that is, the true value of  $\beta$  is zero), show that if  $0 < \alpha_0 < 1/3$  then

$$\sqrt{T} S_\beta(\theta_0) \xrightarrow{d} N(0, \Omega)$$

with

$$\Omega = 8E \left( \frac{x_{t-2}^2}{\omega_0 + \alpha_0 x_{t-1}^2} \right)^2.$$

**Question A.3:**

In the ARCH(2) model we wish to test the hypothesis  $\beta = 0$  by the Quasi likelihood ratio statistic

$$QLR = \left( L(\hat{\theta}) - L(\tilde{\theta}) \right),$$

where  $\tilde{\theta}$  maximizes  $L(\theta)$  with  $\beta = 0$ . The asymptotic distribution of the  $QLR$  statistic is not  $\chi_1^2$ . Explain why it is not  $\chi^2$ . And discuss which distribution it has. Relate your answer to Question A.2.

## Question B:

Consider the model for  $x_t \in \mathbb{R}$  given by

$$x_t = b_{s_t} + a_{s_t}x_{t-1} + \varepsilon_t, \quad (\text{B.1})$$

where the error term  $\varepsilon_t$  satisfies

$$\varepsilon_t \sim i.i.d.N(0, \sigma^2). \quad (\text{B.2})$$

Moreover,

$$b_{s_t} = 1(s_t = 1)b_1 + 1(s_t = 2)b_2, \quad (\text{B.3})$$

and

$$a_{s_t} = 1(s_t = 1)a_1 + 1(s_t = 2)a_2, \quad (\text{B.4})$$

where  $s_t$  is a state variable that takes values in  $\{1, 2\}$  according to the transition probabilities

$$P(s_t = j | s_{t-1} = i) = p_{ij}, \quad (\text{B.5})$$

and  $1(s_t = i) = 1$  if  $s_t = i$  and  $1(s_t = i) = 0$  if  $s_t \neq i$  for  $i = 1, 2$ . We assume throughout that the processes  $(\varepsilon_t)$  and  $(s_t)$  are independent. The model parameters satisfy  $b_1, b_2, a_1, a_2 \in \mathbb{R}$  and  $\sigma^2 > 0$ .

**Question B.1:** When is the process  $(s_t)$  weakly mixing?

**Question B.2:** Let  $f(x_t | x_{t-1}, s_t)$  denote the conditional density of  $x_t$  given  $(x_{t-1}, s_t)$ . Provide an expression for  $f(x_t | x_{t-1}, s_t)$ .

**Question B.3:** In the following we assume that  $p_{11} = 1 - p_{22} =: p \in (0, 1)$  such that  $(s_t)$  is an *i.i.d.* process with  $P(s_t = 1) = p$ . This implies that  $f(x_t | x_{t-1}) = f(x_t | x_{t-1}, s_t = 1)P(s_t = 1) + f(x_t | x_{t-1}, s_t = 2)P(s_t = 2) > 0$ . Use this to show that  $x_t$  satisfies the drift criterion with drift function  $\delta(x) = 1 + x^2$  when

$$a_1^2 p + a_2^2 (1 - p) < 1.$$

**Question B.4:** Consider the *restricted* version of the model in (B.1)-(B.4), where  $b_1 = b_2 = 0$ . Maintaining the assumptions from Question B.3, and

assuming that  $(s_t)$  is *observed*, we consider the log-likelihood function (up to a constant),

$$L_T(\theta) = \sum_{t=1}^T \left[ 1(s_t = 1) \left\{ -\frac{1}{2} \log(\sigma^2) - \frac{(x_t - a_1 x_{t-1})^2}{2\sigma^2} + \log(p) \right\} + 1(s_t = 2) \left\{ -\frac{1}{2} \log(\sigma^2) - \frac{(x_t - a_2 x_{t-1})^2}{2\sigma^2} + \log(1-p) \right\} \right],$$

where  $\theta = (p, a_1, a_2, \sigma^2)$ . Show that the Maximum Likelihood Estimator for  $a_1$  is

$$\hat{a}_1 = \frac{\sum_{t=1}^T 1(s_t = 1) x_t x_{t-1}}{\sum_{t=1}^T 1(s_t = 1) x_{t-1}^2}.$$

Assume that the joint process  $(s_t, x_{t-1})$  is weakly mixing such that  $E[x_{t-1}^2] < \infty$ . Argue that  $\hat{a}_1 \xrightarrow{p} a_1$  as  $T \rightarrow \infty$ .

**Question B.5:** Suppose that the process  $(s_t)$  is *unobserved*, but still satisfies the *i.i.d.* assumption, i.e.  $p_{11} = (1 - p_{22}) = p \in (0, 1)$ . Then the estimator  $\hat{a}_1$  derived in Question B.4 is infeasible. Consider instead the function

$$L_T^\dagger(\theta) = E[L_T(\theta) | x_1, \dots, x_T].$$

It holds (still with  $b_1 = b_2 = 0$ ) that

$$L_T^\dagger(\theta) = \sum_{t=1}^T \left[ P_t^\dagger(1) \left\{ -\frac{1}{2} \log(\sigma^2) - \frac{(x_t - a_1 x_{t-1})^2}{2\sigma^2} + \log(p) \right\} + (1 - P_t^\dagger(1)) \left\{ -\frac{1}{2} \log(\sigma^2) - \frac{(x_t - a_2 x_{t-1})^2}{2\sigma^2} + \log(1-p) \right\} \right].$$

where  $P_t^\dagger(1) = P(s_t = 1 | x_t)$ .

Explain briefly the role of  $L_T^\dagger(\theta)$  for the estimation of  $\theta$ .